

The sample median

Assume that the observations are ranked from the smallest to the largest as:

$$x_{(1)}, x_{(2)}, \dots, x_{(n)}$$

The sample median, \tilde{x} , is then the observation in the middle.

$$\tilde{x} = \begin{cases} x_{\left(\frac{n+1}{2}\right)}, & \text{if } n \text{ is odd} \\ \frac{1}{2}(x_{\left(\frac{n}{2}\right)} + x_{\left(\frac{n}{2}+1\right)}), & \text{if } n \text{ is even} \end{cases}$$

The sample median is more robust against outliers than the average. It represents the value for which 50% of the observations are smaller and is also called the 2. sample quartile, Q_2 . The first quartile, Q_1 , represents the value for which 25% of the observations are smaller.

Similarly Q_3 (3. quartile) represents the value for which 75% of the observations are smaller.

The way MINITAB finds Q_1 and Q_3 is:

Q_1 is the value in position $\frac{n+1}{4}$

Q_3 is the value in position $3\left(\frac{n+1}{4}\right)$

In the case of 10 observations

$$\frac{m+1}{4} = 2.75, \quad 3\left(\frac{m+1}{4}\right) = 8.25$$

$$\Rightarrow Q_1 = \frac{1}{4} x_{(2)} + \frac{3}{4} x_{(3)} \text{ and } Q_3 = \frac{3}{4} x_{(8)} + \frac{1}{4} x_{(9)}$$

Box-plot

Compute Q_2 = sample median

compute $h = Q_3 - Q_1$ = the height of the box

Observations further away than $\frac{3}{2}h$ from the top or
the bottom of the box are called outliers.

Those further away than $3h$ are called strong outliers.

A line is drawn from the top and the bottom of
the box to the largest and the smallest observation that
is not an outlier.

Normal-plot

Let $X \sim N(\mu, \sigma^2)$

$$F_x(x) = P(X \leq x) = P\left(\frac{X-\mu}{\sigma} \leq \frac{x-\mu}{\sigma}\right) = \Phi\left(\frac{x-\mu}{\sigma}\right)^{N(0,1)}$$

$$\Rightarrow \Phi^{-1}(F_x(x)) = \frac{x-\mu}{\sigma}$$

Let x_1, x_2, \dots, x_n be a random sample from a
normal distribution. Ordered by size the values
(realizations) can be written as

$x_{(1)}, x_{(2)}, \dots, x_{(m)}$

Then $F_X(x_{(i)}) \approx \frac{\#\{x_i \leq x_{(i)}\}}{m} = \frac{i}{m} \approx \begin{cases} \frac{i - \frac{1}{2}}{m} \\ \frac{i - \frac{3}{8}}{m + \frac{1}{4}} \end{cases} = f_i$

A plot of $\bar{\Phi}'(f_i)$ against $x_{(i)}$ should approximately become a straight line

With method A Inference from the use of method A

$x_1, \dots, x_{10} \sim N(\mu_A, \sigma_A^2)$, μ_A and σ_A is unknown.

Is the expected mean less than 85?

$$H_0: \mu_A = 85 \quad H_1: \mu_A < 85$$

The random sample approach

Test statistic $\frac{\bar{X} - 85}{\frac{\sigma_A}{\sqrt{m}}} \sim N(0, 1)$, but since σ_A is unknown

where $m = 10$.

we must use $T = \frac{\bar{X} - 85}{\frac{s_A}{\sqrt{m}}} \sim t_9$ Chapter 9.4, 9.5, 10.4

$$\text{Here } s_A = \sqrt{\frac{1}{9} \sum_{i=1}^{10} (x_i - \bar{x})^2}$$

$$t_{\text{obs}} = \frac{84.24 - 85}{\frac{2.902}{\sqrt{10}}} = -0.83. \quad P(T \leq -0.83 | H_0) = 0.214$$

No particular reason to reject H_0 .

Want to find out if $\mu_B > \mu_A$

$$H_0: \mu_B = \mu_A \quad H_1: \mu_B > \mu_A \quad \text{chapter (10)}$$

The random sample approach

If $X_1, \dots, X_{10} \sim N(\mu_A, \sigma_A^2)$ and independent and $Y_1, \dots, Y_{10} \sim N(\mu_B, \sigma_B^2)$ and independent and independent of X_1, \dots, X_{10} , then

$$\frac{\bar{Y} - \bar{X} - (\mu_B - \mu_A)}{\sqrt{\frac{\sigma_B^2}{10} + \frac{\sigma_A^2}{10}}} \sim N(0, 1)$$

Since $\text{Var}(\bar{Y}) = \frac{\sigma_B^2}{10}$ and $\text{Var}(\bar{X}) = \frac{\sigma_A^2}{10}$ chapter (8.4, 9.8, 10.5)

But σ_A^2 and σ_B^2 are unknown

$$S_A^2 = \frac{1}{9} \sum_{i=1}^{10} (X_i - \bar{X})^2, \quad S_B^2 = \frac{1}{9} \sum_{i=1}^{10} (Y_i - \bar{Y})^2 \quad \text{chapter (8)}$$

If $\sigma_A^2 = \sigma_B^2 = \sigma^2$ we have

$$S_p^2 = \frac{\sum_{i=1}^{10} (X_i - \bar{X})^2 + \sum_{i=1}^{10} (Y_i - \bar{Y})^2}{10 + 10 - 2} \quad \text{Kap (9.8, 10.5)}$$

and $T = \frac{\bar{Y} - \bar{X} - (\mu_B - \mu_A)}{S \sqrt{\frac{1}{10} + \frac{1}{10}}}$ is t_{18}

Under $H_0: \mu_B - \mu_A = 0$ and $t_{\text{obs}} = \frac{85.54 - 84.24}{3.3 \sqrt{\frac{1}{10} + \frac{1}{10}}} = 0.88$

$$P(T \geq 0.88) = 0.195$$

No particular reason to claim that $\mu_B > \mu_A$

$$T_B \quad \sigma_A^2 = \sigma_B^2 ?$$

If $\sigma_A^2 \neq \sigma_B^2$ we have

$$T = \frac{\bar{Y} - \bar{X} - (\mu_B - \mu_A)}{\sqrt{\frac{\sigma_A^2}{10} + \frac{\sigma_B^2}{10}}} \sim t\text{-distributed with}$$

$$V = \frac{\left(\frac{\sigma_A^2}{10} + \frac{\sigma_B^2}{10}\right)^2}{\left(\frac{\sigma_A^2}{10}\right)^2/9 + \left(\frac{\sigma_B^2}{10}\right)^2/9}$$

$$t_{obs} = \frac{85.54 - 84.24}{\sqrt{\frac{(2.4)^2}{10} + \frac{(3.65)^2}{10}}} = 0.88$$

$$V = 17 \Rightarrow P(T \geq t_{obs} | H_0) = 0.195$$

An alternative approach

Use of historical data

For method A we have 210 data: x_1, \dots, x_{210}

I: A: x_1, \dots, x_{10} B: x_{11}, \dots, x_{20}

J: A: x_{11}, \dots, x_{10} B: x_{12}, \dots, x_{21}

K: A: x_{191}, \dots, x_{200} B: x_{201}, \dots, x_{210}

Only 9 out of 191 differences are greater than 1.3

$P(\bar{X}_B - \bar{X}_A > 1.3) = \frac{9}{191} = 0.047 < 0.05$ and may indicate that Method B is better than Method A